

## Lecture 20:

### More Complex Models

Background:

In the previous arrival model, we assume the arrival rate is a constant  $\lambda$ . In a more complicated case, the arrival rate may vary on time, i.e.,  $\lambda = \lambda(t)$ . Moreover, different students might have various arrival probabilities. In this lecture, assume the total number of students is  $n$  and for each  $1 \leq m \leq n$ , the arrival probability is  $p_{n,m}$ . Thus, one can use  $X_{n,m} \sim \text{Bernoulli}(p_{n,m}) = \text{binomial}(1, p_{n,m})$  to represent the arrival event for  $m$ -th student of the total  $n$ -student case.

Definition 20.1. (Total Variation norm and Total Variation distance)

Suppose that  $\mu$  is a signed measure on  $(\Omega, \Sigma)$ .

$$\|\mu\|_{TV} := \sup\{\mu(A) \mid A \in \Sigma\} - \inf\{\mu(A) \mid A \in \Sigma\}.$$

Suppose that  $X$  &  $Y$  are integer-valued random variables,

$$d_{TV}(X, Y) = \max_{A \subseteq \mathbb{Z}} |\mathbb{P}(X \in A) - \mathbb{P}(Y \in A)|.$$

Remark 20.1.  $d_{TV}(X, Y) = \|X - Y\|_{TV} = \frac{1}{2} \sum_k |\mathbb{P}(X=k) - \mathbb{P}(Y=k)|$   
 $\|X - Y\|_1$ :  $L_1$  distance between  $X$  and  $Y$ .

Remark 20.2. One can check that  $d_{TV}$  is a metric

①.  $d_{TV}(X, Y) \geq 0$ ; and "=" holds iff  $X=Y$ ;

②.  $d_{TV}(X, Y) = d_{TV}(Y, X)$ ;

③.  $d_{TV}(X, Y) \leq d_{TV}(X, Z) + d_{TV}(Z, Y)$ .

Definition 20.2. (Kullback-Leibler divergence)

Suppose  $\mu$  and  $\nu$  are two probability distributions defined on the space  $\mathcal{X}$ , then

$$D_{KL}(\mu \parallel \nu) = \sum_{x \in \mathcal{X}} \mu(x) \cdot \log \frac{\mu(x)}{\nu(x)}.$$

Remark 20.3. (Pinsker's Inequality).  $\|\mu - \nu\|_{TV} \leq \sqrt{\frac{1}{2} D_{KL}(\mu \parallel \nu)}$ .

Definition 20.3. (Convergence in TV norm).

We say random variables  $X_n \rightarrow X$  in TV if

$$\lim_{n \rightarrow \infty} \|X_n - X\|_{TV} = 0.$$

Theorem 20.1. Suppose  $X_{n,m} \sim \text{Bernoulli}(p_{n,m})$  are independent.

Let  $S_n = \sum_{m=1}^n X_{n,m}$ ,  $\lambda_n = \mathbb{E}S_n = \sum_{m=1}^n p_{n,m}$ , and

$Y_n = \text{Poisson}(\lambda_n)$ . Then

$$\|S_n - Y_n\|_{TV} \leq \sum_{m=1}^n p_{n,m}^2.$$

If, moreover,  $\sup_m p_{n,m} \xrightarrow{n \rightarrow \infty} 0$ ,  $\lambda_n \xrightarrow{n \rightarrow \infty} \lambda$ ,

and  $Y = \text{Poisson}(\lambda)$ , then

$$\lim_{n \rightarrow \infty} \|S_n - Y\|_{TV} = 0,$$

i.e.,  $S_n \rightarrow \text{Poisson}(\lambda)$  in TV norm.

Proof,

$$\begin{aligned} & \|S_n - Y_n\|_{TV} \\ &= \left\| \sum_{m=1}^n X_{n,m} - \sum_{m=1}^n \text{Poisson}(p_{n,m}) \right\|_{TV} \\ &= \left\| \sum_{m=1}^n (X_{n,m} - \text{Poisson}(p_{n,m})) \right\|_{TV} \end{aligned}$$

$$\leq \sum_{m=1}^n \|X_{n,m} - \text{Poisson}(p_{n,m})\|_{TV}$$

$$= \sum_{m=1}^n \frac{1}{2} \sum_{k=0}^{\infty} |\mathbb{P}(X_{n,m} = k) - \mathbb{P}(\text{Poisson}(p_{n,m}) = k)|$$

$$= \frac{1}{2} \sum_{m=1}^n \left\{ |(1-p_{n,m}) - e^{-p_{n,m}}| + |p_{n,m} - p_{n,m} e^{-p_{n,m}}| + \sum_{k=2}^{\infty} \mathbb{P}(\text{Poisson}(p_{n,m}) = k) \right\}$$

$$1 \geq e^{-p_{n,m}} \geq 1 - p_{n,m}$$

$$= \frac{1}{2} \sum_{m=1}^n \left\{ e^{-p_{n,m}} - 1 + p_{n,m} + p_{n,m} - p_{n,m} e^{-p_{n,m}} + 1 - e^{-p_{n,m}} - p_{n,m} e^{-p_{n,m}} \right\}$$

$$= \sum_{m=1}^n p_{n,m} (1 - e^{-p_{n,m}})$$

$$\leq \sum_{m=1}^n p_{n,m}^2$$

Pinsker's inequality

why?

$$\log x \leq x - 1$$

Notice that  $\|Y_n - Y\|_{TV}$

$$\leq \sqrt{\frac{1}{2} D_{KL}(Y_n \| Y)}$$

$$= \sqrt{\frac{1}{2} (\lambda - \lambda_n + \lambda_n \cdot \log \frac{\lambda_n}{\lambda})}$$

$$\leq \sqrt{\frac{1}{2} (\lambda - \lambda_n + \lambda_n \cdot \frac{\lambda_n - \lambda}{\lambda})}$$

$$= \frac{|\lambda_n - \lambda|}{\sqrt{2\lambda}}$$

Thus, by the triangular inequality,

$$0 \leq \|S_n - Y\|_{TV}$$

$$\leq \|S_n - Y_n\|_{TV} + \|Y_n - Y\|_{TV}$$

$$\leq \sum_{m=1}^n p_{n,m}^2 + \frac{|\lambda_n - \lambda|}{\sqrt{2\lambda}}$$

$$\leq \left( \max_{1 \leq m \leq n} p_{n,m} \right) \cdot \left( \sum_{m=1}^n p_{n,m} \right) + \frac{|\lambda_n - \lambda|}{\sqrt{2\lambda}}$$

$$= \left( \max_{1 \leq m \leq n} p_{n,m} \right) \cdot \lambda_n + \frac{|\lambda_n - \lambda|}{\sqrt{2\lambda}}.$$

Since  $\lim_{n \rightarrow \infty} \left( \max_{1 \leq m \leq n} p_{n,m} \right) = 0$  and  $\lim_{n \rightarrow \infty} \lambda_n = \lambda$ ,

one has  $\lim_{n \rightarrow \infty} \left( \max_{1 \leq m \leq n} p_{n,m} \right) \cdot \lambda_n + \frac{|\lambda_n - \lambda|}{\sqrt{2\lambda}} = 0$ .

By the squeeze theorem, we know

$$\lim_{n \rightarrow \infty} \|S_n - Y\|_{TV} = 0. \quad \square$$

Example 20.1. Let  $p_{n,m} = \frac{\lambda}{n}$ ,  $\lambda_n = \lambda$ , thus

$$\max_m p_{n,m} = \frac{\lambda}{n} \xrightarrow{n \rightarrow \infty} 0 \quad \text{and} \quad \lambda_n \xrightarrow{n \rightarrow \infty} \lambda.$$

By Theorem 20.1,

$$S_n = X_{n,1} + X_{n,2} + \cdots + X_{n,n} \xrightarrow{TV} \text{Poisson}(\lambda).$$

This implies Theorem 18.1. Thus, one can

view Theorem 18.1 as a special case of Theorem 20.1.

Example 20.2. Suppose  $p_{n,m}$  is not a constant but

$$\max_{1 \leq m \leq n} p_{n,m} \xrightarrow{n \rightarrow \infty} 0 \quad \text{and} \quad \lambda_n \xrightarrow{n \rightarrow \infty} \lambda. \quad \text{Then}$$

$$S_n \xrightarrow{TV} \text{Poisson}(\lambda).$$

Example 20.3. Let  $\lambda: [0,1] \rightarrow \mathbb{R}_+$  be continuous and

$$p_{n,m} = \frac{1}{n} \lambda\left(\frac{m}{n}\right), \quad \forall n \in \mathbb{N}, \quad \forall 1 \leq m \leq n. \quad \text{Then}$$

$$0 \leq \max_{1 \leq m \leq n} p_{n,m} \leq \frac{1}{n} \max_{t \in [0,1]} \lambda(t) \xrightarrow{n \rightarrow \infty} 0,$$

$$\text{and} \quad \lambda_n = \sum_{m=1}^n \frac{1}{n} \lambda\left(\frac{m}{n}\right) \xrightarrow{n \rightarrow \infty} \int_0^1 \lambda(t) dt. \quad \text{Thus,}$$

$$S_n \longrightarrow \text{Poisson}\left(\int_0^1 \lambda(t) dt\right).$$

Definition 20.4. (Non-homogeneous Poisson Process)

Let  $N(t)$  represents the total number of occurrence or events that have happened up to and including time  $t$ . We say that

$\{N(s), s \geq 0\}$  is a non-homogeneous Poisson Process with rate  $\lambda(r)$  if

(i).  $N(0) = 0$ ,

(ii).  $N(t+s) - N(s) = \text{Poisson}(\int_s^{t+s} \lambda(r) dr)$ , and

(iii).  $N(t)$  has independent increments.

Remark 20.4. Interarrivals are in general neither exponential distributions nor independent. For instance,

$$\begin{aligned} \mathbb{P}(\tau_1 > t) &= \mathbb{P}(N(t) = 0) = \mathbb{P}(\text{Poisson}(\int_0^t \lambda(r) dr) = 0) \\ &= e^{-\int_0^t \lambda(r) dr} \end{aligned}$$

is not exponential unless  $\lambda$  is a constant.

This is the end of this lecture !